



# A Class of Purely Sequential Minimum Risk Point Estimation Methodologies with Second-Order Properties

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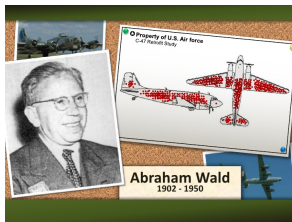
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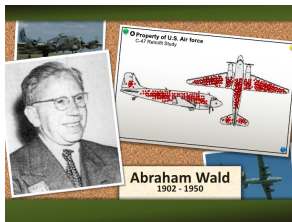
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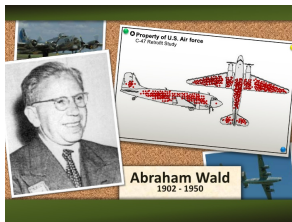


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- ▶ The sample size is not predetermined.
- ▶ One observation is recorded at a time successively until termination.



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- ▶ Loss function:

$$L_n \equiv L_n(\mu, \bar{X}_n) = A(\bar{X}_n - \mu)^2 + cn, \quad (1)$$

where  $A(> 0)$  and  $c(> 0)$  are both known.





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- ▶ Risk function:

$$R_n(c) \equiv E_{\mu, \sigma}[L_n(\mu, \bar{X}_n)] = A\sigma^2 n^{-1} + cn. \quad (2)$$



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- ▶ NO fixed-sample-size procedure.



## 1.2. Minimum Risk Point Estimation (MRPE)

# Solutions

- ▶ Two-stage: Stein (1945,1949)
- ▶ Purely sequential: Robbins (1959), Starr (1966)
- ▶ Three-stage: Mukhopadhyay (1990)
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### 1.3. Estimators for $\sigma$

- ▶  $\sigma$  is unknown.
- ▶ A general arbitrary estimator, assumed positive w.p.1.,

$$W_n \equiv W_n(X_1, \dots, X_n).$$

1.3. Estimators for  $\sigma$ Conditions on  $W_n$ 

C1 *Independence*:  $\bar{X}_n$  and  $\{W_k; 2 \leq k \leq n\}$  are distributed independently for all  $n \geq 2$ .



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- C3 *Asymptotic normality*:  $\sqrt{n}(\sigma^{-1}W_n - 1) \xrightarrow{\mathcal{L}} N(0, \delta^2)$  as  $n \rightarrow \infty$ .

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- C3 *Asymptotic normality*:  $\sqrt{n}(\sigma^{-1}W_n - 1) \xrightarrow{\mathcal{L}} N(0, \delta^2)$  as  $n \rightarrow \infty$ .
- C4 *Uniform continuity in probability*: For every  $\varepsilon > 0$ , there exists a large  $\nu$  and small  $\gamma > 0$  for which  $\forall n \geq \nu$ ,

$$P_{\mu, \sigma} \left( \max_{0 \leq k \leq n\gamma} |W_{n+k} - W_n| \geq \varepsilon \right) < \varepsilon.$$

1.3. Estimators for  $\sigma$ 

**C5** *Kolmogorov's inequality*: For every  $\varepsilon > 0$ , and some  $2 \leq n_1 \leq n_2$ , with  $r \geq 2$ ,

$$P_{\mu, \sigma} \left( \max_{n_1 \leq n \leq n_2} |W_n - \sigma| \geq \varepsilon \right) \leq \varepsilon^{-r} E_{\mu, \sigma} [|W_{n_1} - \sigma|^r].$$

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**C7** *Wiener's condition*:  $E_{\mu,\sigma}[\sup_{n \geq 2} W_n] < \infty$ .



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$$\mathcal{P} : N_{\mathcal{P}} \equiv N_{\mathcal{P}}(c) = \inf\{n \geq m(\geq 2) : n \geq \sqrt{A/c}(W_n + n^{-\lambda})\}, \quad (5)$$

where  $\lambda(> \frac{1}{2})$  is held fixed.





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where  $\lambda(> \frac{1}{2})$  is held fixed.

- ▶  $P_{\mu,\sigma}\{N_{\mathcal{P}} < \infty\} = 1$  and  $N_{\mathcal{P}} \uparrow \infty$  w.p.1 as  $c \downarrow 0$ .



## 2.1 Methodologies

► Upon  $\{N_{\mathcal{P}}, X_1, \dots, X_m, X_{m+1}, \dots, X_{N_{\mathcal{P}}}\}$ :

$$\bar{X}_{N_{\mathcal{P}}} \equiv N_{\mathcal{P}}^{-1} \sum_{j=1}^{N_{\mathcal{P}}} X_j. \quad (6)$$



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► *Risk Efficiency:*

$$\xi_{\mathcal{P}}(c) = \frac{R_{N_{\mathcal{P}}}(c)}{R_{n^*}(c)} = \frac{1}{2}E_{\mu,\sigma}[N_{\mathcal{P}}/n^*] + \frac{1}{2}E_{\mu,\sigma}[n^*/N_{\mathcal{P}}];$$



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► *Regret:*

$$\omega_{\mathcal{P}}(c) = R_{N_{\mathcal{P}}}(c) - R_{n^*}(c) = cE_{\mu,\sigma}[(N_{\mathcal{P}} - n^*)^2/N_{\mathcal{P}}].$$



## 2.2 Asymptotics

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### ▶ Asymptotic Second-Order Risk Efficiency:

$$\omega_{\mathcal{P}}(c) = \delta^2 c + o(c) \text{ as } c \rightarrow 0, \quad (9)$$

with  $\delta^2$  coming from (C3).



## Illustrations

- ▶ What kinds of  $W_n$  would satisfy (C1)-(C7)?





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- ▶ Consider  $W_n$  what involves only

$$\mathbf{Y}_n = (X_1 - X_n, X_2 - X_n, \dots, X_{n-1} - X_n).$$



## Illustration 0: Sample Standard Deviation

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- ▶ The regret expansion:

$$\delta^2 = \frac{1}{2} \Rightarrow \omega_{\mathcal{P}_0}(c) = \frac{1}{2}c + o(c).$$



## Illustration 1: Gini's Mean Difference (GMD)

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$$\text{GMD: } g_n = \binom{n}{2}^{-1} \sum \sum_{1 \leq i < j \leq n} |X_i - X_j|.$$



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- ▶ The regret expansion:

$$\delta^2 = \frac{\pi + 6\sqrt{3} - 12}{3} \approx 0.511 \Rightarrow \omega_{\mathcal{P}_1}(c) = 0.511c + o(c).$$





## Illustration 2: Mean Absolute Deviation (MAD)

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- ▶ The regret expansion:

$$\delta^2 = \frac{\pi - 2}{2} \approx 0.571 \Rightarrow \omega_{\mathcal{P}_2}(c) = 0.571c + o(c).$$



**Table 1.** Simulations from  $N(5, 4)$  with  $A = 100, m = 10, \lambda = 2$   
under 1000 runs implementing  $\mathcal{P}_0 - \mathcal{P}_2$

$n^*$	$100c$	$\mathcal{P}$	$\bar{n}$	$s(\bar{n})$	$\hat{\xi}$	$s(\hat{\xi})$	$\delta^2$	$\hat{\omega}/c$
50	16	$\mathcal{P}_0$	50.012	0.1671	0.9880	0.003340	0.5	0.593131
		$\mathcal{P}_1$	50.313	0.1703	0.9879	0.003377	0.511	0.612431
		$\mathcal{P}_2$	50.259	0.1778	0.9872	0.003339	0.571	0.666431
100	4	$\mathcal{P}_0$	99.955	0.2408	0.9932	0.002404	0.5	0.599650
		$\mathcal{P}_1$	100.335	0.2347	0.9943	0.002306	0.511	0.561200
		$\mathcal{P}_2$	100.332	0.2495	0.9939	0.002327	0.571	0.636125
200	1	$\mathcal{P}_0$	200.012	0.3325	0.9969	0.001660	0.5	0.561100
		$\mathcal{P}_1$	200.255	0.3380	0.9971	0.001661	0.511	0.580400
		$\mathcal{P}_2$	200.026	0.3562	0.9962	0.001659	0.571	0.643800
400	0.25	$\mathcal{P}_0$	399.931	0.4588	0.9983	0.001146	0.5	0.531200
		$\mathcal{P}_1$	400.282	0.4508	0.9984	0.001114	0.511	0.514000
		$\mathcal{P}_2$	400.232	0.4873	0.9985	0.001145	0.571	0.598800



## Accelerated Sequential MRPE Saving Sampling Operations

- ▶ Given the pilot sample size  $m \geq 2$ ,  $0 < \rho \leq 1$  and  $k \geq 1$ , an integer, consider the following stopping rule:

$$T \equiv T(c) = \inf \left\{ n \geq 0 : m + kn \geq \rho \sqrt{A/c} \left[ W_{m+kn} + (m + kn)^{-\lambda} \right] \right\}.$$



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- ▶ Operational time reduced by approximately  $100(1 - k^{-1}\rho)\%$ .





## Selected References

- ▶ Gini, C. (1914). *L'Ammontare la Composizione della Ricchezza delle Nazioni*, Boca, Torino.
- ▶ Gini, C. (1921). Measurement of Inequality of Incomes, *Economic Journal* 31: 124-126.
- ▶ Hu, J. and Mukhopadhyay, N. (2019) Second-Order Asymptotics in a Class of Purely Sequential Minimum Risk Point Estimation (MRPE) Methodologies, *Japanese Journal of Statistics and Data Science* 2: 81-104.
- ▶ Mukhopadhyay, N. (1996). An Alternative Formulation of Accelerated Sequential Procedures with Applications to Parametric and Nonparametric Estimation, *Sequential Analysis* 15: 253-269.



- ▶ Mukhopadhyay, N., and Hu, J. (2017). Confidence Intervals and Point Estimators for a Normal Mean Under Purely Sequential Strategies Involving Gini's Mean Difference and Mean Absolute Deviation, *Sequential Analysis* 36: 210-239.
- ▶ Mukhopadhyay, N. and Solanky, T. K. S. (1991). Second Order Properties of Accelerated Stopping Times with Applications in Sequential Estimation, *Sequential Analysis* 10: 99-123.
- ▶ Robbins, H. (1959). Sequential Estimation of the Mean of a Normal Population, in *Probability and Statistics*, H. Cramér volume, Ulf Grenander, ed., pp. 235-245, Uppsala: Almqvist & Wiksell.
- ▶ Starr, N. (1966). On the Asymptotic Efficiency of a Sequential Procedure for Estimating the Mean, *Annals of Mathematical Statistics* 37: 1173-1185.



- ▶ Stein, C. (1945). A Two Sample Test for a Linear Hypothesis Whose Power Is Independent of the Variance, *Annals of Mathematical Statistics* 16: 243-258.
- ▶ Stein, C. (1949). Some Problems in Sequential Estimation (Abstract), *Econometrica* 17: 77-78.



Thank  
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